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Factors associated with exclusive breastfeeding at hospital discharge: a study using data from the Georgian Birth Registry



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Abstract

Background: The World Health Organization recommends exclusive breastfeeding for six months, defined as no other solids or liquids besides breast milk and essential vitamins or medicines. Data about exclusive breastfeeding are limited in Georgia, and the information that exist are provided by national surveys, that present inconsistent numbers. Georgia has recently established a national birth registry, which includes information about early postpartum breastfeeding. The objective of this study was to identify factors associated with exclusive breastfeeding of term newborns at hospital discharge in Georgia, using national registry data.

Methods: All live, singleton, term births registered in the Georgian Birth Registry in November and December 2017 were included, with a final study sample of 7134 newborns. Newborns exclusively breastfed at hospital discharge were compared with those who were not, and potential factors were assessed with logistic regression analysis. Hospital discharge normally occurred between 2 and 5 days postpartum.

Results: The study identified several factors associated with nonexclusive breastfeeding of term newborns at hospital discharge in Georgia: maternal higher education compared to secondary education or less (Adjusted Odds Ratio [AOR] 0.75; 95% CI 0.59, 0.97), caesarean delivery compared to vaginal or assisted vaginal delivery (AOR 0.47; 95% CI 0.37, 0.60), birthweight < 2500 g compared to 3000–3499 g (AOR 0.51; 95% CI 0.27, 0.97), and admission to neonatal intensive care unit after delivery (AOR 0.02; 95% CI 0.02, 0.03). None of the following factors were associated with exclusive breastfeeding at discharge: mother's age, marital status, Body Mass Index (BMI), parity, in vitro fertilization, maternal intrapartum complications and the sex of the newborn.

Conclusions: To the authors' knowledge, this is the first time determinants of exclusive breastfeeding at hospital discharge have been studied in Georgia. Several factors associated with nonexclusive breastfeeding at discharge were identified, most noteworthy were caesarean delivery and admission to neonatal intensive care unit. These findings are of importance to the Georgian health authorities and maternal/child non-governmental organizations.

Keywords: Exclusive breastfeeding, Patient discharge, Caesarean section, Neonatal intensive care units, Social determinants of health, Georgia (republic)

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Background

The World Health Organization (WHO) recommends exclusive breastfeeding for the first six months of a child's life, defined as no other solids or liquids besides breast milk and essential vitamins or medicines, and continued breastfeeding up to the age of two years or beyond [1]. To facilitate exclusive breastfeeding, the WHO promotes immediate skin-to-skin contact between mothers and newborns, and initiation of breastfeeding within one hour of delivery [2]. In low- and middleincome countries, only 37% of children under the age of six months are exclusively breastfed [3], despite the fact that breastfeeding is associated with numerous shortand long-term health advantages for both the mother and child. For mothers, long-term benefits include lower risk of breast cancer, ovarian cancer, and type II diabetes mellitus [4]. In children, breastfeeding has a protective effect against gastrointestinal and respiratory infections below five years of age [5], with a clear dose-response relationship [6] of long-term effects include lower odds of obesity and overweight in childhood and later life [7, 8].

Many factors affect breastfeeding; they can be maternal, newborn, or obstetric in nature and are often interlinked. Older age of the mother has been positively associated with exclusive breastfeeding (EBF) at discharge [9]. The relation between maternal age and EBF might be partly explained by parity, but this link is uncertain. Overweight and obesity has been associated with lower odds of exclusive breastfeeding [10], which may be related to both physical and psychosocial factors. Higher education and high socio-economic status of the mother are associated with increased rates of breastfeeding initiation and duration in high-income countries, for both exclusive and any breastfeeding [10-14], whereas in lowand middle-income countries this association is inversed [15–19]. Studies from all over the world indicate a negative association between caesarean delivery and breastfeeding [9, 14, 15, 20-25]. In addition to maternal and/ or newborn distress, the effect of caesarean delivery on early breastfeeding might be related to delayed onset of lactation, problems with newborn suckling, disrupted early skin-to-skin contact and mother-newborn interaction, and postoperative hospital practices [23].

Located in the Caucasus region, Georgia is categorized as an upper-middle income developing country, ranked 70 of 188 countries in the 2017 Human Development Index [26]. Georgia has a population of 3.7 million; 87% consider themselves as ethnic Georgians and almost 11% are from the neighboring countries Armenia and Azerbaijan [27]. The infant mortality rate decreased from 22.5 in 2009 to 9.6 per 1000 livebirths in 2017, with a fertility rate of 2.1 in 2017 [27]. All medical facilities in Georgia are private. The beneficiaries of private insurance have decreased since Georgia introduced Universal Health Care in 2013. Basic obstetric care is free, which currently includes eight antenatal visits and the delivery. All Georgian mothers deliver at maternity wards with qualified healthcare personnel (99.9%) [28], and the average length of stay for vaginal delivery is 3–4 days and for caesarean delivery 5–6 days [29].

During the early 2000s, awareness increased about the advantages of breastfeeding in Georgia. In 2004, 14 out of 78 maternity wards in the country were designated as baby-friendly through the Baby-Friendly Hospital Initiative (BFHI), a global initiative to promote and support breastfeeding by WHO and the United Nations Children's Fund [30]. However, today there are no babyfriendly hospitals left in Georgia due to lack of follow-up on the initiative and breastfeeding support after discharge is lacking at primary healthcare level. Data on breastfeeding practice in Georgia are limited. The information that does exist are provided by national surveys, which have presented inconsistent numbers. One survey from 2010 showed that 20% of mothers initiated breastfeeding within one hour of delivery [29], while another survey from the same year reported a proportion of 66% [31]. According to a survey from 2005, only 11% of children under six months were exclusively breastfed [32], whereas in 2010 this proportion was reported to be as much as 55% [31].

Georgia introduced a national birth registry in 2016, the Georgian Birth Registry (GBR) [33], which collects medical data from antenatal visits, as well as data from the delivery and postpartum period until discharge from the maternity ward/hospital. Reporting to the GBR is mandatory. Designated health professionals receive training in how to manage the register, and are then responsible for teaching the staff at their own facility. The GBR includes information on early postpartum breastfeeding. The objective of this study was to identify factors associated with exclusive breastfeeding of term newborns at hospital discharge in Georgia, using national registry data.

Methods

The study population consisted of all births registered in the GBR from November 1st to December 31st 2017 (n = 8159 newborns). Because this study looks at term newborns only (born between gestational age 37^{+0} and 41^{+6} weeks) [34], all preterm and post-term births were excluded (n = 721). Gestational age was estimated by the first day of the last menstrual period in 70.1% of the sample, with the rest being estimated by ultrasound. Stillbirths (n = 76) and neonatal deaths (n = 38) were also excluded, as were newborns with a HIV-positive mother (n = 5) or a surrogate mother (n = 52). Multiple births (n = 246) were excluded because these mothers face greater challenges to initiate and to continue

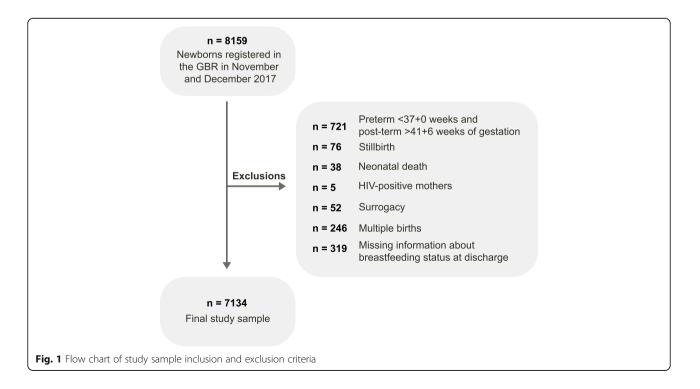
Variables

The outcome variable of interest was exclusive breastfeeding at hospital discharge, defined as the newborn receiving only breast milk the last feedings before discharge. Nonexclusive breastfeeding included documentation of mixed feeding, formula feeding, or parenteral feeding (intravenous administration of nutrition) in the last feedings before hospital discharge. Newborn discharge normally occurred between two to five days postpartum, with a median of three days for vaginal delivery and four days for caesarean delivery. Potential factors were selected based on the existing literature. The variables had to be available in the GBR dataset and reliable for the study purpose. Three categories of variables were considered: maternal, delivery, and newborn.

The maternal variables were mother's age, marital status, education level, body mass index (BMI), parity and in vitro fertilization. Mother's age at delivery was given in years and used in a continuous form in the regression analysis. For descriptive purposes, mother's age was further categorized into five groups: < 20, 20–24, 25–29, 30-34, and ≥ 35 years. The youngest mother in the study sample was 14 years and the oldest was 52, which is

biologically plausible; thus, no cut-offs were applied. Marital status was categorized as single, which included a small number of divorced women (n = 2), married, and unknown. Education level was categorized as completed secondary education or less, higher education, and unknown. BMI was computed by dividing the mother's weight in kg at the first antenatal visit (before week 12 of pregnancy) by her height in m^2 . If the weight was not measured before week 12, the mother's self-reported pre-pregnancy weight was used. BMI was then categorized in accordance with the WHO classification system [36]: underweight ($< 18.5 \text{ kg/m}^2$), normal weight (18.5 - <25 kg/m²), overweight (25- < 30 kg/m²) and obese (≥ 30 kg/m²). Extreme BMI values (< 15 and \ge 60 kg/m²) were excluded from the statistical analysis [37]. Parity (i.e., the total number of previous live and stillbirths per woman), was categorized as 0, 1, 2, and ≥ 3 births. Three or more births were merged into one group because of low numbers. In vitro fertilization for the current birth was included as a dichotomous variable.

The delivery variables included were mode of delivery and maternal intrapartum complications. Mode of delivery was dichotomized as caesarean delivery (including both elective and emergency caesarean delivery) and vaginal delivery (including assisted delivery with forceps or vacuum, manual handling of breech delivery, and cases of episiotomy). Maternal intrapartum complications that may affect breastfeeding were constructed as *one* dichotomous variable and coded as "yes" if the mother experienced any of the following conditions: placenta praevia,



placental abruption, meconium in amniotic fluid, umbilical cord prolapse, shoulder dystocia, uterine rupture, eclampsia during labor, retained placenta, uterine atony, or hemorrhage with total amount of bleeding > 500 ml.

Newborn variables were admission to neonatal intensive care unit (NICU) after delivery, birth weight and newborn sex. A NICU admission right after delivery leads to separation of mother and newborn, disrupting the early initiation of breastfeeding. Admission to NICU was included as a dichotomous variable. Unknown NICU admission was coded as missing due to the small numbers in the sample (n = 129). Birthweight was measured in grams and categorized as < 2500 g, 2500–2999 g, 3000–3499 g, 3500–3999 g, and ≥ 4000 g. Extreme values of < 500 g and > 7000 g were excluded from the analysis. The group that included the mean value (3000–3499 g) was set as the reference.

Statistical analysis

The statistical analysis was conducted using R version 3.4.3 [38]. Demographic characteristics of the mothers and their newborns according to breastfeeding status at hospital discharge were compared using independent ttest, Wilcoxon rank-sum test, or chi-square test when appropriate. Logistic regression was used to assess potential factors associated with the newborn being exclusively breastfed at hospital discharge. The analyses were performed in accordance with the model building strategy described in Veierød et al. [39]. Variables that were significant at the 0.25 level in the univariable model were initially selected for inclusion in the multivariable model. Stepwise elimination was applied, and the full and reduced models were compared using the likelihood ratio test if more than one variable was removed. The changes in the coefficients (ß) of the remaining variables were computed to test if the removed variable(s) were needed to adjust for others in the model. If one coefficient changed by more than 20%, the variable was kept as a confounder in the final model. These steps were repeated until all variables left in the model were either significant or important confounders of other variables in the model. Only subjects without missing values on the variables in the final model were included (n = 6993). The following plausible interactions between the variables in the model were tested: parity and maternal age, parity and education level, parity and delivery mode, education level and delivery mode, maternal age and delivery mode, and NICU and birthweight. None of the interactions were significant. Case-wise diagnostics and multicollinearity were examined. The final model was tested for overall goodness-of-fit using the Hosmer-Lemeshow test. The results are presented with odds ratios (OR) and confidence intervals (CI), using the significance level 0.05.

Results

Of the 7134 newborns in the study sample, 6583 (92.3%) were exclusively breastfed at hospital discharge and 551 were not. Maternal, delivery, and newborn characteristics differed by breastfeeding status at hospital discharge (Table 1).

In adjusted analyses, education level, mode of delivery, birthweight, and admission to NICU were identified as factors associated with nonexclusive breastfeeding at hospital discharge. Mothers with higher education were 25% less likely to exclusively breastfeed at hospital discharge compared to mothers with secondary education or less (AOR 0.75; 95% CI 0.59, 0.97). Newborns born by caesarean delivery were 53% less likely to be exclusively breastfed at hospital discharge compared to those born by vaginal delivery (AOR 0.47; 95% CI 0.37, 0.60). The newborns with the lowest birth weight (< 2500 g)had 49% lower odds of receiving exclusive breastfeeding at hospital discharge compared to those weighing 3000-3499 g (AOR 0.51; 95% CI 0.27, 0.97). In addition, admission to NICU after delivery was strongly associated with nonexclusive breastfeeding at hospital discharge. Newborns admitted to NICU were 98% less likely to be exclusively breastfed at hospital discharge compared to newborns who were not admitted (AOR 0.02; 95% CI 0.02, 0.03). None of the tested interaction terms were significant (Table 2).

Discussion

The study identified several factors associated with nonexclusive breastfeeding at hospital discharge: maternal higher education compared to secondary education or less, caesarean delivery compared to vaginal delivery, birth weight < 2500 g compared to 3000–3499 g, and admission to NICU.

In line with research from other low- and middleincome countries [15-19], the current study showed that mothers with higher education were less likely to exclusively breastfeed their newborns at hospital discharge. This finding is the opposite of what is seen in most high-income countries [10–14]. The association between education and exclusive breastfeeding in Georgia might be related to the work situation of mothers and the prospect of paid maternity leave. The unemployment rate among women in Georgia is low (11.2% in 2018, which is 2.7% lower than men) [40]. According to the Labour code of Georgia, a mother is entitled to 183 days (~26 weeks) of paid maternity leave financed by the state with a maximum of 1000 Georgian Lari [41], or around 305 Euro for the whole maternity leave. The legislation also states that employers and employees can agree on additional benefits [41]. However, for higher educated mothers with a well-paid job and for those without extra benefits, the compensation may be insufficient. If the

	Exclusive breastfeeding at discharge	Nonexclusive breastfeeding at discharge	p - value	Both groups combined	Total N^b
N (mother-newborn pairs)	6583	551		7134	
Mothers					
Age at delivery in years mean (SD)	27.4 (5.63)	28.4 (5.75)	< 0.001	27.5 (5.65)	
Age at delivery in years %			0.003		
< 20	6.8	5.3		6.7	475
20–24	26.4	21.4		26.0	1855
25–29	33.2	33.9		33.3	2373
30–34	21.8	22.9		21.9	1560
≥ 35	11.8	16.5		12.2	871
Marital status %			0.13		
Single	13.1	11.3		13.0	924
Married	50.1	54.4		50.4	3598
Unknown	36.8	34.3		36.6	2612
Education level %			< 0.001		
Secondary school or less	56.7	48.5		56.0	3996
Higher education	34.9	45.6		35.8	2551
Unknown	8.4	6.0		8.2	586
BMI in kg/m ² , median (25th–75th percentile)	22.9 (20.5–26.0)	23.2 (20.8–26.6)	0.14	22.9 (20.6–26.1)	
BMI in kg/m², %			0.23		
< 18.5	7.5	7.0		7.5	465
18.5- < 25	60.3	59.5		60.3	3735
25- < 30	21.6	20.0		21.5	1334
≥ 30	10.5	13.4		10.7	663
Parity %			0.03		
0	39.8	43.0		40.1	2860
1	38.4	32.5		37.9	2707
2	16.9	18.0		17.0	1210
≥ 3	4.9	6.5		5.0	357
In vitro fertilization %	0.7	0.9	0.42	0.7	48
Delivery					
Mode of delivery %			< 0.001		
Vaginal delivery	58.3	46.2		57.3	4089
Caesarean delivery	41.7	53.8		42.7	3041
Maternal intrapartum complications %	4.8	5.4	0.58	4.9	347
Newborns					
Newborn sex %			0.88		
Female	48.3	48.7		48.3	3447
Male	51.7	51.3		51.7	3685
Birthweight in g, mean (SD)	3364 (426.5)	3274 (516.9)	< 0.001	3357 (434.8)	
Birthweight in g %			< 0.001		
< 2500	1.5	5.1		1.8	125
2500–2999	15.4	19.2		15.7	1121

 Table 1
 Maternal, delivery, and newborn characteristics according to breastfeeding status at hospital discharge^a

	Exclusive breastfeeding at discharge	Nonexclusive breastfeeding at discharge	p - value	Both groups combined	Total N^b
N (mother-newborn pairs)	6583	551		7134	
3000-3499	44.0	41.9		43.9	3127
3500-3999	30.5	26.1		30.1	2148
≥ 4000	8.6	7.6		8.5	606
Admission to NICU %	1.8	43.7	< 0.001	4.4	308

Table 1 Maternal, delivery, and newborn characteristics according to breastfeeding status at hospital discharge^a (Continued)

^aThe Georgian Birth Registry, November–December 2017 (n = 7134)

SD standard deviation, BMI body mass index, NICU neonatal intensive care unit

^bFor some variables, the numbers do not add up to the total (n = 7134) because of missing values: education n = 7133, BMI n = 6197, mode of delivery n = 7130, newborn sex n = 7132, birthweight n = 7127, and admission to NICU n = 7005

mother returns to work early for economical or other reasons, exclusive breastfeeding in the postpartum period would probably not be a top priority.

Newborns born by caesarean delivery were less likely to be exclusively breastfed at hospital discharge compared to newborns born by vaginal or assisted vaginal delivery. This finding is in accordance with other studies [9, 21–23]. The proportion of caesarean deliveries is noticeably higher in Georgia (43.5% in 2016) [33] compared to the mean in European countries (25.2% in 2010) [42]. Considering this high rate, the negative association with early postpartum exclusive breastfeeding is of particular interest. Indeed, excessive rates of caesarean delivery is associated with several short- and long-term health risks for both the mother and newborn [43], and lower odds of exclusive breastfeeding at hospital discharge adds to the list of health risks. One of the Ten steps to Successful Breastfeeding in the BFHI is immediate skin-to-skin contact and early initiation of breastfeeding [44]. Without the designation of baby-friendly hospitals, Georgian hospitals and maternity wards may not pay enough attention to the early initiation of breastfeeding after a caesarean delivery. One systematic review suggested that adequate breastfeeding support after a caesarean delivery reduces the negative association between caesarean delivery and early initiation of breastfeeding entirely [45]. Another review found that, among mothers that successfully initiated breastfeeding, there was no difference in exclusive breastfeeding between babies delivered by caesarean and those born by vaginal delivery at six months [23]. These findings indicate that a supportive breastfeeding environment after a caesarean delivery, with proper postoperative pain management, could substantially improve the rates of exclusive breastfeeding at hospital discharge and later in the postpartum period. The Georgian Ministry of Health is addressing the high caesarean rates, and hospitals are under scrutiny to reduce their rates. However, it is too soon to evaluate the results of this intervention.

Newborn admission to a NICU had a large negative impact on exclusive breastfeeding at hospital discharge,

a finding that is in line with other studies [46, 47]. This is an expected finding, as exclusive breastfeeding often cannot be prioritized in an intensive care setting. However, although the present study excluded premature newborns, the association between NICU admission and exclusive breastfeeding at hospital discharge was very strong (AOR 0.02; 95% CI 0.02, 0.03). Additionally, newborns with a birthweight < 2500 g were less likely to be exclusively breastfed at hospital discharge compared to newborns weighing 3000-3499 g. Low birthweight has been associated with lower odds of exclusive breastfeeding in previous studies as well [48]. An expansion of the BFHI has been developed for use in NICUs, where the Ten Steps to Successful Breastfeeding are adapted for preterm and sick newborns. The steps include early and prolonged mother-newborn skin-to-skin contact, also known as Kangaroo Mother Care, and support of early breastfeeding, with newborn physiological stability as the only criterion,- not newborn age, weight, or other criteria [49]. This is important, as early initiation of breastfeeding reduces the risk of neonatal mortality, also among low-birthweight newborns [50]. The BFHI expansion to NICUs recommends rooming-in, where mothers and newborns stay together in the NICU, as a step to facilitate continuous breastfeeding [49]. Rooming-in at NICUs may not be feasible in all settings, but the mother should then get the opportunity to stay close to the NICU.

The absolute majority of the term newborns in the study were exclusively breastfed at discharge (92.3%). The high proportion raises concern of reporting bias. Nurses and midwives should record newborn feeding twice a day, but we cannot be sure that all healthcare personnel follow the same procedures, with the risk of misclassification by type of newborn feeding at discharge. However, before applying the exclusion criteria, 85.0% of all newborns were exclusively breastfed at hospital discharge. Prevalence figures of exclusive breastfeeding at hospital discharge in other countries vary substantially, from 61.6% in Canada (only term newborns) [9], 82.7% in rural Western Australia [47], 86.9

Table 2 Odds ratios with 95% confidence intervals of exclusive breastfeeding at hospital discharge^a

	Univariable analysis ^b	Multivariable analysis	
Mother's age at delivery in years	0.97 [0.96, 0.98]	0.98 [0.96, 1.00]	
Marital status			
Single	1.26 [0.96, 1.69]	_	
Married	Reference	-	
Unknown	1.17 [0.97, 1.41]	_	
Education level		-	
Secondary school or less	Reference	Reference	
Higher education	0.66 [0.55, 0.79]	0.75 [0.59, 0.97]	
Unknown	1.20 [0.84, 1.77]	1.28 [0.82, 2.09]	
BMI in kg/m ²			
< 18.5	1.06 [0.74, 1.56]	-	
18.5- < 25	Reference	-	
25- < 30	1.06 [0.84, 1.36]	-	
≥ 30	0.77 [0.58, 1.03]	-	
Parity			
0	Reference	Reference	
1	1.28 [1.04, 1.56]	1.13 [0.86, 1.47]	
2	1.01 [0.80, 1.30]	0.89 [0.64, 1.26]	
≥3	0.81 [0.56, 1.18]	0.81 [0.49, 1.37]	
In vitro fertilization for current birth	0.72 [0.31, 2.08]	_	
Mode of delivery			
Vaginal delivery	Reference	Reference	
Caesarean delivery	0.61 [0.52, 0.73]	0.47 [0.37, 0.60]	
Maternal intrapartum complications	0.88 [0.61, 1.32]	_	
Newborn sex			
Female	Reference	-	
Male	1.02 [0.85, 1.21]	_	
Birthweight in g			
< 2500	0.28 [0.18, .44]	0.51 [0.27, 0.97]	
2500–2999	0.76 [0.60, 0.97]	0.83 [0.61, 1.14]	
3000–3499	Reference	Reference	
3500–3999	1.11 [0.90, 1.38]	1.17 [0.88, 1.55]	
≥ 4000	1.07 [0.77, 1.53]	1.18 [0.78, 1.86]	
Admission to NICU	0.02 [0.02, 0.03]	0.02 [0.02, 0.03]	

^aThe Georgian Birth Registry, November–December 2017 (n = 7134)

BMI body mass index, NICU neonatal intensive care unit

^bComplete case analysis: education level n = 7133, BMI n = 6197, mode of delivery n = 7130, newborn sex n = 7132, birthweight n = 7127, and admission to NICU n = 7005

^cComplete case analysis: *n* = 6993

-93.1% in the Czech Republic [51], to 93.5% in rural China (only healthy singletons) [52]. The prevalence figures from other countries support the plausibility of the high proportion of exclusive breastfeeding we observed in Georgia. Although it is previously reported a lower prevalence of breastfeeding in Georgia [29, 31, 32], the data are 10–15 years old and based on surveys.

Furthermore, the data are not directly comparable as the surveys display the initiation and duration of exclusive breastfeeding. The rate of exclusive breastfeeding is likely much lower at 4–6 months, because the majority of mothers have returned to work or studies. A pilot follow-up program (up to 6 months) in the GBR was launched in 2019 in the region of Ajara in West Georgia.

When this program is expanded, we will be able to conduct a follow-up study on the duration and potential obstacles of breastfeeding after hospital discharge.

The GBR was launched in 2016, and it takes time to establish good practices for proper reporting [33]. This is evident for some variables, for instance maternal intrapartum complications. Even though the variable merged several complications the mother may experience during delivery, the prevalence of these complications was low (total 4.9%), indicating a considerable underreporting of complications in the GBR. This underreporting can mask a potential effect of maternal intrapartum complications on exclusive breastfeeding at hospital discharge. Additionally, data are entered into the GBR by maternity wards all over the country, indicating that the healthcare personnel have different training in how to assess certain maternal and newborn conditions. The variable maternal intrapartum complications are particularly exposed, because many of the complications included in the variable are based on the judgement of the healthcare personnel attending the birth. Registry data is also vulnerable to data entry errors. Both of these problems could potentially introduce bias, but we do not know to what extent. Data on the validity of the variables in the GBR are still not available, hence the findings should be interpreted with some caution.

Strengths and limitations

The GBR collects information on the absolute majority of births in Georgia, and participation in the registry is mandatory by law. The study population was limited to the two last months of 2017 because of structural changes in GBR variables, which were completed in late October of 2017. Selected variables (mother's age, education level, BMI and mode of delivery) were compared between the study population before exclusions and total births in 2017. Based on these comparisons, there are no significant differences (< 1.5%) between the study sample and total births in 2017, thus the study sample is representative. Data from 2018 were not available at the time of submission.

Even though the study comprised a relatively large sample size (n = 7134), the number of newborns who were not exclusively breastfed at hospital discharge was small (n = 551) compared to those who were (n = 6583). This may have led to larger CIs than if the groups had been more equal in size or if the sample size had been bigger. The study adjusted for several confounders, but not for gestational age. Even though the study population only consisted of newborns at term, important differences between gestational age of 37^{+0} weeks and 41^{+6} weeks could still exist. In particular, gestational age is closely interlinked with the newborn factors birthweight and admission to NICU, and including gestational age as

a confounder would likely have rendered these estimates more conservative than the presented results.

Several studies suggest a difference in breastfeeding outcomes between emergency and elective caesarean delivery [9, 23, 53]. However, these subgroups were not included in the study because of possible misclassification in favor of emergency deliveries in the GBR. In the initial population, 30.8% of all births were emergency caesarean deliveries and 14.3% were elective caesarean deliveries (data not shown). In comparison, the median proportion of emergency caesarean deliveries was 12.9% and for elective caesarean deliveries 10.7% in an aggregated study of European countries [42]. Had reliable data been available, separating the caesarean deliveries into elective and emergency would be important, as it would allow the adjustment for related variables like in vitro fertilization, maternal intrapartum complications, and birthweight.

Of the 8159 newborns in the initial study population, 135 were excluded due to missing or unknown breastfeeding status at hospital discharge. Compared to the study sample (n = 7134), these excluded newborns experienced a higher proportion of caesarean delivery and admission to NICU, both of which were factors associated with nonexclusive breastfeeding at hospital discharge. This suggests that there was a higher proportion of newborns who were not exclusively breastfed at discharge in the excluded cases. Hence, missing information on the main outcome variable could have led to an underestimation of the ORs for caesarean delivery and admission to NICU. Some of the included predictor variables had missing data; however, all variables had less than 0.1% missing, except for BMI with 13.1% missing. The cases with missing maternal BMI were compared with the included cases, and they did not differ significantly. Thus, the assumption is that the exclusion of these cases did not bias the effect estimates.

Conclusions

To the authors' knowledge, this is the first time exclusive breastfeeding at hospital discharge have been studied in Georgia, and certainly for the first time using national birth registry data. The study identified several factors associated with nonexclusive breastfeeding at hospital discharge in Georgia in term newborns: maternal higher education, caesarean delivery, low birthweight and admission to NICU after delivery. Potential steps to increase rates of exclusive breastfeeding are a reintroduction of the Baby-Friendly Hospital Initiative, reducing the high caesarean delivery rates in Georgia, as well as strengthening the length and economical support of the maternity leave. These findings will be valuable for national health authorities when setting new priorities in maternal and child health, as well as for nongovernmental organizations working with breastfeeding dyads. Hopefully, the findings will increase awareness about breastfeeding in maternity wards and hospitals all over Georgia. The results could have validity in countries with high rates of caesarean delivery and similar health system structures as in Georgia.

Abbreviations

BFHI: Baby-Friendly Hospital Initiative; GBR: Georgian Birth Registry; HIV: Human immunodeficiency virus; NICU: Neonatal Intensive Care Unit; WHO: World Health Organization

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Authors' contributions

MSL wrote the manuscript and performed the statistical analyses. IHN and EEA supervised the statistical analyses and revised the manuscript. All authors read and approved the final manuscript.

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Availability of data and materials

The National Centre for Disease Control and Public Health (NCDC), Tbilisi, provides data storage, administration, and quality control of the GBR. The datasets generated and analyzed during the current study are not publicly available without a written application and acceptance by NCDC.

Ethics approval and consent to participate

Submission of data to the GBR is mandatory by law for pregnant women in Georgia. The dataset was completely anonymized before analysis. The National Center for Disease Control and Public Health Institutional Review Board in Georgia has approved the study - protocol (IRB # 2017–010 31.03.2017). In addition, the Regional Committee for Medical and Health Research Ethics (REC) of Northern Norway concluded that no permission from them is necessary (2017/404/REK Nord). In addition, the legal department at UIT the Arctic University of Norway has concluded that no further permissions from the National Centre for Research Data were needed.

Consent for publication

Not applicable.

Competing interests

The authors declare that they have no competing interests.

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References

- World Health Organization. Global strategy for infant and young child feeding. Geneva: World Health Organization; 2003. http://www.who.int/ nutrition/publications/infantfeeding/9241562218/en/. Accessed 07 Nov 2017.
- World Health Organization. Guideline: protecting, promoting and supporting breastfeeding in facilities providing maternity and newborn services. 2017 [http://www.who.int/nutrition/publications/guidelines/ breastfeeding-facilities-maternity-newborn/en/. Accessed 7 Nov 2017.
- Victora CG, Bahl R, Barros AJD, França GVA, Horton S, Krasevec J, et al. Breastfeeding in the 21st century: epidemiology, mechanisms, and lifelong effect. Lancet. 2016;387(10017):475–90.

- Chowdhury R, Sinha B, Sankar MJ, Taneja S, Bhandari N, Rollins N, et al. Breastfeeding and maternal health outcomes: a systematic review and meta-analysis. Acta Paediatr. 2015;104(467):96–113.
- Horta BL, Victora CG, World Health Organization. Short-term effects of breastfeeding: a systematic review on the benefits of breastfeeding on diarrhoea and pneumonia mortality. Geneva: World Health Organization; 2013. https://apps.who.int/iris/handle/10665/95585. Accessed 7 Nov 2017.
- Kramer MS, Kakuma R. Optimal duration of exclusive breastfeeding. Cochrane Database Syst Rev. 2012;15(8):Cd003517..
- Horta BL, Loret de Mola C, Victora CG. Long-term consequences of breastfeeding on cholesterol, obesity, systolic blood pressure and type 2 diabetes: a systematic review and meta-analysis. Acta Paediatr. 2015; 104(S467):30–7.
- Yan J, Liu L, Zhu Y, Huang G, Wang PP. The association between breastfeeding and childhood obesity: a meta-analysis. BMC Public Health. 2014;14:1267.
- McDonald SD, Pullenayegum E, Chapman B, Vera C, Giglia L, Fusch C, et al. Prevalence and predictors of exclusive breastfeeding at hospital discharge. Obstet Gynecol. 2012;119(6):1171–9.
- Bjorset VK, Helle C, Hillesund ER, Overby NC. Socio-economic status and maternal BMI are associated with duration of breast-feeding of Norwegian infants. Public Health Nutr. 2018:1–9.
- Kristiansen AL, Lande B, Overby NC, Andersen LF. Factors associated with exclusive breast-feeding and breast-feeding in Norway. Public Health Nutr. 2010;13(12):2087–96.
- Kohlhuber M, Rebhan B, Schwegler U, Koletzko B, Fromme H. Breastfeeding rates and duration in Germany: a bavarian cohort study. Br J Nutr. 2008;99: 1127–32.
- van Rossem L, Oenema A, Steegers EA, Moll HA, Jaddoe VW, Hofman A, et al. Are starting and continuing breastfeeding related to educational background? The generation R study. Pediatrics. 2009;123(6):e1017–27.
- Vassilaki M, Chatzi L, Bagkeris E, Papadopoulou E, Karachaliou M, Koutis A, et al. Smoking and caesarean deliveries: major negative predictors for breastfeeding in the mother-child cohort in Crete, Greece (Rhea study). Matern Child Nutr. 2014;10(3):335–46.
- Kasahun AW, Wako WG, Gebere MW, Neima GH. Predictors of exclusive breastfeeding duration among 6-12 month aged children in gurage zone, South Ethiopia: a survival analysis. Int Breastfeed J. 2016;12:20.
- Hossain M, Islam A, Kamarul T, Hossain G. Exclusive breastfeeding practice during first six months of an infant's life in Bangladesh: a country based cross-sectional study. BMC Pediatr. 2018;18(1):93.
- Hazir T, Akram DS, Nisar YB, Kazmi N, Agho KE, Abbasi S, et al. Determinants of suboptimal breast-feeding practices in Pakistan. Public Health Nutr. 2013; 16(4):659–72.
- El-Gilany AH, Shady E, Helal R. Exclusive breastfeeding in Al-Hassa, Saudi Arabia. Breastfeed Med. 2011;6(4):209–13.
- Grjibovski AM, Yngve A, Bygren LO, Sjostrom M. Socio-demographic determinants of initiation and duration of breastfeeding in Northwest Russia. Acta Paediatr. 2005;94(5):588–94.
- Vieira TO, Vieira GO, Giugliani ER, Mendes CM, Martins CC, Silva LR. Determinants of breastfeeding initiation within the first hour of life in a Brazilian population: cross-sectional study. BMC Public Health. 2010;10:760.
- Kling DOI, Haile ZT, Francescon JOI, Chertok I. Association between method of delivery and exclusive breastfeeding at hospital discharge. J Am Osteopath Assoc. 2016;116(7):430–9.
- 22. Liu X, Zhang J, Liu Y, Li Y, Li Z. The association between cesarean delivery on maternal request and method of newborn feeding in China. PLoS One. 2012;7(5):e37336.
- Prior E, Santhakumaran S, Gale C, Philipps LH, Modi N, Hyde MJ. Breastfeeding after cesarean delivery: a systematic review and meta-analysis of world literature. Am J Clin Nutr. 2012;95(5):1113–35.
- Cato K, Sylven SM, Lindback J, Skalkidou A, Rubertsson C. Risk factors for exclusive breastfeeding lasting less than two months - identifying women in need of targeted breastfeeding support. PLoS One. 2017;12(6):e0179402.
- Adugna B, Tadele H, Reta F, Berhan Y. Determinants of exclusive breastfeeding in infants less than six months of age in Hawassa, an urban setting, Ethiopia. Int Breastfeed J. 2017;12:45.
- Human Development Report Office (HDRO), United Nations Development Programme (UNDP). Human development indices and indicators: 2018 statistical update. 2018 [http://www.hdr.undp.org/sites/default/files/2018_ human_development_statistical_update.pdf. Accessed 3 Nov 2017.

- National Statistics Office of Georgia (GeoStat). Statistical yearbook of Georgia. Tbilisi: National Statistics Office of Georgia (GeoStat); 2018. https:// www.geostat.ge/media/20931/Yearbook_2018.pdf. Accessed 30 Oct 2019.
- Gamkrelidze A, Kereselidze M, Tsintsadze M, Gambashidze K, Shakhnazarova M, Tsetskhladze N, et al. Health care - statistical yearbook 2016 - Georgia. Tbilisi: National Centre for Disease Control and Public Health; 2017. http:// www.ncdc.ge/Handlers/GetFile.ashx?ID=31eee2a3-9bf5-4558-959b-a4b92f6 00555. Accessed 16 Mar 2018.
- Division of Reproductive Health CfDCaPC, Georgia Ministry of Labor Health and Social Affairs, National Center for Disease Control and Public Health Georgia, National Statistics Office of Georgia (GeoStat). Reproductive health survey Georgia 2010. 2012 [http://unicef.ge/uploads/1._GERHS_2010__Final_ Report_ENGL_resized.pdf. Accessed 7 Nov 2017.
- Nemsadze K. Report from the country of Georgia: protecting and promoting breastfeeding through regulation of artificial-feeding marketing practices. J Perinat Educ. 2004;13(1):23–8.
- UNICEF Georgia, National Centre for Disease Control and Public Health Georgia. Report of the 2009 Georgia national nutrition survey 2010 [http:// unicef.ge/uploads/Report_of_the_Georgia_National_Nutrition_Survey_2 009_-_eng.pdf. Accessed 3 Nov 2017.
- State Department of Statistics of Georgia, National Centre for Disease Control and Public Health Georgia, UNICEF. Georgia multiple indicator cluster survey 2005. 2008 [https://mics-surveys-prod.s3.amazonaws.com/ MICS3/Europe and Central Asia/Georgia/2005/Final/Georgia 2005 MICS_ English.pdf. Accessed 3 Nov 2017.
- Anda EE, Nedberg IH, Rylander C, Gamkrelidze A, Turdziladze A, Skjeldestad FE, et al. Implementing a birth registry in a developing country-experiences from Georgia. Tidsskrift for den Norske laegeforening: tidsskrift for praktisk medicin, ny raekke. 2017;138(2):138–43.
- World Health Organization. ICD-10: International Statistical Classification of Diseases and Related Health Problems. 10th ed. http://apps.who.int/ classifications/icd10/browse/2016/en. Accessed 13 Mar 2018.
- McAndrew F, Thompson J, Fellows L, Large A, Speed M, Renfrew MJ. Infant feeding survey 2010: NHS information Centre, IFF Research; 2012 [https:// digital.nhs.uk/catalogue/PUB08694. Accessed 27 Feb 2018.
- World Health Organization. BMI classification. [http://apps.who.int/bmi/ index.jsp?introPage=intro_3.html. Accessed 25 Jan 2018.
- The Global BMIMC. Body-mass index and all-cause mortality: individualparticipant-data meta-analysis of 239 prospective studies in four continents. Lancet (London, England). 2016;388(10046):776–86.
- R Core Team. R: A language and environment for statistical computing. R Foundation for Statistical Computing. Vienna: R Foundation for Statistical Computing; 2017.
- Veierød MB, Lydersen S, Laake P. Medical statistics in clinical and epidemiological research: Gyldendal akademisk; 2012.
- National Statistics Office of Georgia (GeoStat). Employment and unemployment. [Internet] GeoStat. 2018. https://www.geostat.ge/en/ modules/categories/38/employment-and-unemployment. Accessed 7 Nov 2019.
- Parliament of Georgia. Labour code of Georgia [Organic Law of Georgia]. [Internet] Parliament of Georgia. 2010. https://matsne.gov.ge/en/document/ view/1155567. Accessed 21 Mar 2018.
- Macfarlane AJ, Blondel B, Mohangoo AD, Cuttini M, Nijhuis J, Novak Z, et al. Wide differences in mode of delivery within Europe: risk-stratified analyses of aggregated routine data from the euro-Peristat study. BJOG Int J Obstet Gynaecol. 2016;123(4):559–68.
- World Health Organization. WHO statement on caesarean section rates. Geneva: World Health Organization; 2015. http://www.who.int/ reproductivehealth/publications/maternal_perinatal_health/cs-statement/ en/. Accessed 20 Nov 2017.
- World Health Organization UNCsF. Baby-friendly hospital initiative revised, updated and expanded for integrated care. 2009 [http://apps.who.int/iris/ bitstream/10665/43593/1/9789241594967_eng.pdf. Accessed 6 Mar 2018.
- Rollins NC, Bhandari N, Hajeebhoy N, Horton S, Lutter CK, Martines JC, et al. Why invest, and what it will take to improve breastfeeding practices? Lancet. 2016;387(10017):491–504.
- Davanzo R, Monasta L, Ronfani L, Brovedani P, Demarini S. Breastfeeding at NICU discharge: a multicenter Italian study. J Hum Lact. 2013;29(3):374–80.
- Cox K, Giglia R, Zhao Y, Binns CW. Factors associated with exclusive breastfeeding at hospital discharge in rural Western Australia. J Hum Lact. 2014;30(4):488–97.

- Jones JR, Kogan MD, Singh GK, Dee DL, Grummer-Strawn LM. Factors associated with exclusive breastfeeding in the United States. Pediatrics. 2011;128(6):1117–25.
- Nyqvist KH, Häggkvist A-P, Hansen MN, Kylberg E, Frandsen AL, Maastrup R, et al. Expansion of the baby-friendly hospital initiative ten steps to successful breastfeeding into neonatal intensive care: expert group recommendations. J Hum Lact. 2013;29(3):300–9.
- Smith ER, Hurt L, Chowdhury R, Sinha B, Fawzi W, Edmond KM, et al. Delayed breastfeeding initiation and infant survival: a systematic review and meta-analysis. PLoS One. 2017;12(7):e0180722.
- Mydlilova A, Sipek A, Vignerova J. Breastfeeding rates in baby-friendly and non-baby-friendly hospitals in the Czech Republic from 2000 to 2006. J Hum Lact. 2009;25(1):73–8.
- 52. Tang L, Binns CW, Luo C, Zhong Z, Lee AH. Determinants of breastfeeding at discharge in rural China. Asia Pac J Clin Nutr. 2013;22(3):443–8.
- Hobbs AJ, Mannion CA, McDonald SW, Brockway M, Tough SC. The impact of caesarean section on breastfeeding initiation, duration and difficulties in the first four months postpartum. BMC Pregnancy Childbirth. 2016;16:90.

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